

## Estimation of bigeye catches by purse seiners



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## 1 Introduction

The purse seine fishery in the western and central Pacific Ocean is largely composed of vessels from Japan, Korea, Taiwan and the United States. These fleets catch a variety of tunas, such as skipjack (Katsuwonus pelamis), yellowfin (Thunnus albacares), and bigeye (Thunnus obesus). Reported catch estimates from logsheets are biased for yellowfin and bigeye because bigeye are rarely distinguished from yellowfin. This results in an overestimate of yellowfin and a corresponding under-estimate of bigeye in the fishery.

Bigeye catch estimates are determined for the Japanese and US fleets by port sampling. Estimates for the remaining fleets are based on the assumption that the proportions of bigeye in associated and unassociated sets are similar to those for the US fleet, though the procedure may be biased because bigeye proportions in unassociated and associated sets conducted by the other fleets may differ due to various factors (e.g. time and space variability, fleet behavior, gear modifications).

Estimates of yellowfin and bigeye catches in the purse seine fishery have been modified by the NMFS and SPC to account for misidentification (Coan et al. 2001, Lawson 2000). Modifications are largely dependent on a comprehensive NMFS port sampling programme that monitors purse seine landings in Pago Pago, American Samoa where $85 \%$ of the US catch was landed in 2000 (Coan et al. 2001). Through port sampling, bigeye catches for the entire US fishery are produced by estimating the proportion of bigeye in the combined catches of yellowfin and bigeye. Annual estimates for the remaining fleets are produced by applying the proportion of bigeye in the US catch (described in Lawson 2000). The proportions are weighted by school type because the composition of yellowfin and bigeye is higher in associated purse seine sets (e.g. drifting rafts, logs and FADs) than in unassociated or free swimming sets.

A tree-based regression method was presented at SCTB12 (Bigelow 1999) as an alternative to extrapolation to estimate purse-seine bigeye catch.

The objectives of this paper are to:

1. Briefly compare the composition of bigeye in the yellowfin+bigeye estimates from port and at-sea observer sampling programmes.
2. Consider other models (additive and delta) to examine factors affecting the composition of bigeye in the purse-seine fishery.
3. Provide an update of bigeye catches using the tree-based regression approach that considers all available port and at-sea observer data.

## 2 Data

Species composition data by length were available from port and observer sampling (Table 1). Length data were converted into weight $\left(\mathrm{W}(\mathrm{kg})=8.6388 \mathrm{E}-06 * \mathrm{~L}(\mathrm{~cm})^{3.2174}\right)$. A total of 5,840 samples could be related to other additional factors such as time (year and month), location ( $5^{\circ}$ square), set type and vessel flag. There were 4,599 observer and 1,241 port samples. Observer data were compiled from the USMLT, SPC and Micronesian Fisheries Authority (MFA) programs. Overall, the US fleet was the most sampled ( $\mathrm{n}=2,627$ ).

The types of purse seine sets sampled by port samplers and observers were: 1) free swimming school, 2) feeding on baitfish, 3) drifting log/debris, 4) drifting FAD, 5) anchored FAD, 6) whale, 7) whale shark and 8) boat. These set types were reduced into three set types for further statistical analysis: 1) unassociated (free swimming school and feeding on baitfish), 2)
drifting log and 3) FAD (drifting, anchored, whale, whale shark and boat). In general, the spatial distribution of samples from the individual fleets followed the distribution of actual fleet effort (Figures 1-4). One exception is the Philippines domestic purse-seine fishery where no species composition samples were available.

Prediction models for bigeye composition were constructed using variables in the portsampling and observer data that were similar to variables in the corresponding logbook data in order to estimate catches. Variables considered were: 1) year, 2) month, 3) set type ( $1=$ unassociated, $2=\log$ and $3=$ FAD sets), 4) fleet, 5) latitude and 6 ) longitude.

## 3 Comparison of bigeye composition estimates from port and atsea observer sampling programmes

The annual proportions of bigeye in the bigeye+yellowfin estimates were compared for port and observer sampling (Figures 5-6). For unassociated sets, the bigeye composition was relatively low compared to associated sets (cf. Figures $5 \& 6$ ). From observer sampling, the median estimates were essentially zero (Figure 5A); however, bigeye comprised 5-10\% of the bigeye+yellowfin composition in the initial years of the US port-sampling programme.

For associated sets (drifting logs and FADs), there were differences between the annual estimates based on sampling programmes for a given fleet (Figure 6). The composition of bigeye was consistently higher in port samples compared to observer samples for both the Japanese (Figure 5A \& B) and US fleets (Figure 5C \& D). The differences in annual estimates between sampling programmes could be related to vessel differences, spatial or temporal effects or reflect the high variability in bigeye composition in associated sets. There is a need to further compare the two sampling programmes for sources of bias.

## 4 Models for analyzing factors affecting the composition of bigeye and predicting total catch

### 4.1 Geneneralized additive (GAMs) and delta models

A generalized additive model (GAM) was fit to the six variables to determine the relationship between explanatory variables and bigeye composition (Mean bigeye composition in bigeye+yellowfin=f(year, month, set type, fleet, latitude, longitude) assuming a Poisson distribution. Four of the variables (year, month, set type, fleet) were considered as categorical, whereas latitude and longitude were considered as continuous. The continuous variables were modeled by a loess smoothing function.

All variables in the GAM were considered highly significant from a stepwise GAM process. The relative effects of the explanatory variables on bigeye composition is illustrated in Figure 7. The relative magnitude of the explanatory variables can be inferred from the relative $y$-axis ranges. The variable latitude has the largest effect ( $y$-axis range of 4 ) and suggests that bigeye composition is low at high latitudes such as New Zealand. Though latitude has the largest effect it is relatively unimportant because there are few samples outside the traditional latitudinal range of the fishery $\left(5^{\circ} \mathrm{N}-15^{\circ} \mathrm{S}\right)$.

The effects on bigeye composition are intermediate for set type, year and flag. As expected, unassociated sets have significantly less bigeye than log and FAD sets. The year effect indicates that bigeye proportion has been declining continuously since 1996 when all other variables are considered in the model. The effect of flag is lowest for the Solomon Islands fleet, but could result from a small sample size.

The effects on bigeye composition are smallest for month and longitude. The effect of month is varied without trend. The effect of longitude indicates that bigeye composition is low in the western Pacific $\left(\sim 130^{\circ}\right.$ E), increases steadily throughout the western Pacific, peaks at the dateline $\left(180^{\circ}\right)$ and declines again in the central Pacific from $175^{\circ}$ to $155^{\circ} \mathrm{W}$.

Preliminary delta models were also fit to the port and observer sampling data. The model is fit in a two stage process, with the first model fit to the entire data set with the bigeye composition as presence or absence and a second model that includes only samples where the bigeye composition was greater than zero. The first model assumed a binomial distribution whereas the second model used a normal, lognormal or gamma distribution. The overall mean bigeye composition is then the product of the probability of success catch from the binomial model coupled with the model on non-zero catches. The initial results were not encouraging with the delta approach as the binomial model produced unrealistic estimates of the probability of positive bigeye catches.

### 4.2 Tree-based regression

Tree-based regression models are used in classification problems, whereby data are split by binary partitioning into separate subgroups. The model continues to find splits until no further improvement or splitting is necessary. A regression tree was initially grown from the entire set of possible predictors. The tree was examined in terms of its predictors, residual mean deviance, residuals and normal probability plot of residuals. The tree was pruned and sniped through cross-validation procedures to reduce overfitting. Cross-validation indicated a tree with approximately 70 branches (nodes) would be sufficient (Figure 8).

The final tree included all predictor variables. The order of relative importance was 1) set type, 2) year, 3) month, 4) fleet, 5) longitude and 6) latitude. The residuals were not normally distributed in the tree-based regression similar to the GAM, because bigeye proportion data contained a large amount of zero observations. The regression explained slightly greater than $20 \%$ of the variance in the data. Predictions from the tree results were applied to the entire fishery dataset in order to estimate total bigeye catch by fleet. The SPC "Best database" was used in conjunction with the tree results to predict the bigeye proportion using the predictor variables. The total catch was then estimated by multiplying the bigeye proportion by the reported bigeye+yellowfin estimates.

A comparison of bigeye proportions by set types for the US fleet is provided in Table 2. In general, the regression estimates of the bigeye proportion in unassociated sets is consistently higher than the port-sampling data presented in Lawson (2000, proportions provided by NMFS), whereas the proportions in associated sets vary between years.

Total annual bigeye catch estimates for the US fleet are compared from the two different estimation techniques (Table 3). Estimates from 1988 to 1999 were $31 \%$ higher based on the extrapolation of port-sampling data than the regression technique. The discrepancy results from the inclusion of observer data in the regression estimation which has a lower proportion of bigeye than the port-sampling results.

Annual estimates for the western and central Pacific Ocean (WCPO) range from 7 to 36 thousand tonnes based on extrapolation of port-sampling data and 9 to 37 thousand tonnes based on the regression (Figure 9). Contrary to the US fleet estimates, the regression estimates for the entire WCPO were greater than the extrapolation method in all years except 1996. Since 1988, annual regression estimates averaged $20 \%$ greater than estimates made by extrapolation.

The annual predicted proportions of bigeye in the bigeye+yellowfin estimates are presented for each fleet by set type (unassociated, Log \& FAD) in Table 4. In general the bigeye proportion in log and FAD sets ranges from $15-30 \%$ of the total yellowfin+ bigeye.

Using these proportions from the regression technique, the spatial distribution of bigeye catch was calculated for 1999 (Figure 11). The fleets of Japan, Taiwan, Philippines and the US had the highest catches of bigeye. The US fleet in particular has high catches of bigeye given the fleet behavior to conduct FAD associated sets and the distribution near the dateline where the proportion of bigeye appears higher than in the western Pacific.

## Conclusions \& Recommendations

- The Statistics Working Group is invited to note the content of the paper as an alternative method to estimate bigeye catches by purse-seine vessels in the WCPO.
- The GAM and tree-based regressions had similar results on the factors influencing the proportion of bigeye in purse-seine catches.
- The tree-based method has an advantage over the extrapolation from port-sampling by statistically incorporating differences in fleet behaviour and temporal and spatial variability which are not inherent in the estimates by extrapolation.
- The discrepancy between the port-sampling and observer estimates requires further attention. Bigeye composition in associated set types is characterised by high variability between sets; however, there may be continued identification difficulties in the observer sampling that leads to a downward bias or factors in the port-sampling that represent an upward bias.


## 5 Acknowledgements

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## 6 References

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Table 1. Number of bigeye composition samples from the purse-seine fishery (1988-2000).

| Fleet | Number of observer samples |  | Number of port samples |  |
| :--- | :---: | :---: | :---: | :---: |
|  | Unassociated | Associated | Unassociated | Associated |
| FSM | 35 | 95 |  |  |
| Japan | 85 | 227 | 74 | 246 |
| Korea | 396 | 241 |  |  |
| PNG | 26 | 86 |  |  |
| Philippines | 2 | 225 |  |  |
| Solomon Is. | 4 | 51 |  |  |
| Taiwan | 598 | 745 | 112 | 809 |
| USA | 246 | 1460 |  |  |
| Vanuatu | 45 | 32 | 186 | 1055 |
| Total | 1437 | 3162 |  |  |

Table 2. Comparison of bigeye proportions by set type for the US purse seine fishery based on portsampling data (Lawson 2000) and tree-based regression.

|  | US Port sampling (Lawson 2000) <br> Proportion of bigeye in YFT+BET <br> Year |  | Tree-based estimates  <br> Proportion of bigeye in YFT+BET  <br> Unassociated  |  |
| :---: | :---: | :---: | :---: | :---: |
| 1988 | 0.030 | 0.089 | 0.050 | Associated |

Table 3. Comparison of annual catches (metric tonnes) of bigeye by purse seiners in the WCPO based on port-sampling data (Lawson 2000) and tree-based regression.

|  |  | US fleet |  | All fleets |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Port-sampling | Regression | \% difference | Port-sampling | Regression | \% difference |
| 1988 | 1,948 | 1,078 | 81 | 7,305 | 8,921 | -18 |
| 1989 | 2,421 | 2,888 | -16 | 12,651 | 16,640 | -24 |
| 1990 | 1,762 | 3,522 | -50 | 12,143 | 24,221 | -50 |
| 1991 | 1,550 | 2,748 | -44 | 13,406 | 23,391 | -43 |
| 1992 | 3,480 | 3,479 | 0 | 19,384 | 19,610 | -1 |
| 1993 | 3,731 | 3,539 | 5 | 14,286 | 16,267 | -12 |
| 1994 | 1,711 | 3,101 | -45 | 11,178 | 21,848 | -49 |
| 1995 | 3,190 | 2,413 | 32 | 14,222 | 16,046 | -11 |
| 1996 | 9,860 | 3,591 | 175 | 18,244 | 17,945 | 2 |
| 1997 | 10,058 | 13 | 3,875 | 41 | 31,637 | 36,901 |

Table 4. Proportions of bigeye tuna in the yellowfin+bigeye by set type (unassociated, log and FAD) and fleet estimated by tree-based regression.

| Year | $\begin{gathered} \text { Fleet } \\ \text { AU } \end{gathered}$ |  |  |  | Fleet ES |  |  |  | Fleet |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled |
| 1988 | 0.037 | 0.336 |  | 0.236 |  |  |  |  |  |  |  |  |
| 1989 | 0.037 | 0.325 |  | 0.152 |  |  |  |  |  |  |  |  |
| 1990 | 0.045 | 0.271 | 0.261 | 0.171 |  |  |  |  |  |  |  |  |
| 1991 | 0.347 | 0.283 | 0.449 | 0.299 |  |  |  |  | 0.037 | 0.091 | 0.091 | 0.076 |
| 1992 | 0.035 | 0.278 | 0.111 | 0.187 |  |  |  |  | 0.036 | 0.102 | 0.096 | 0.074 |
| 1993 | 0.027 | 0.449 | 0.449 | 0.164 |  |  |  |  | 0.030 | 0.098 | 0.093 | 0.071 |
| 1994 |  |  |  |  |  |  |  |  | 0.033 | 0.161 | 0.172 | 0.103 |
| 1995 |  |  |  |  |  |  |  |  | 0.036 | 0.173 | 0.111 | 0.091 |
| 1996 |  |  |  |  |  |  |  |  | 0.020 | 0.154 | 0.144 | 0.137 |
| 1997 |  |  |  |  |  |  |  |  | 0.035 | 0.182 | 0.121 | 0.117 |
| 1998 |  |  |  |  |  |  |  |  | 0.033 | 0.110 | 0.123 | 0.057 |
| 1999 |  |  |  |  | 0.037 |  | 0.202 | 0.201 | 0.032 | 0.115 | 0.139 | 0.107 |
| Avg | 0.088 | 0.324 | 0.317 | 0.201 | 0.037 |  | 0.202 | 0.201 | 0.032 | 0.132 | 0.121 | 0.093 |


|  | Fleet |  |  |  | Fleet |  |  |  | Fleet |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | ID |  |  |  | JP |  |  | KI |  |  |  |  |  |
| Year | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled |  |
| 1988 | 0.085 | 0.098 | 0.100 | 0.096 | 0.040 | 0.100 | 0.100 | 0.092 |  |  |  |  |  |
| 1989 | 0.098 | 0.131 | 0.169 | 0.125 | 0.038 | 0.139 | 0.117 | 0.107 |  |  |  |  |  |
| 1990 | 0.086 | 0.211 | 0.231 | 0.150 | 0.042 | 0.188 | 0.156 | 0.112 |  |  |  |  |  |
| 1991 | 0.071 | 0.158 | 0.154 | 0.114 | 0.039 | 0.170 | 0.131 | 0.094 |  |  |  |  |  |
| 1992 | 0.102 | 0.100 | 0.100 | 0.101 | 0.051 | 0.099 | 0.100 | 0.072 |  |  |  |  |  |
| 1993 | 0.088 | 0.097 | 0.099 | 0.092 | 0.032 | 0.098 | 0.099 | 0.059 |  |  |  |  |  |
| 1994 | 0.093 | 0.171 | 0.172 | 0.123 | 0.031 | 0.195 | 0.202 | 0.110 | 0.017 | 0.276 |  | 0.273 |  |
| 1995 | 0.107 | 0.123 | 0.133 | 0.116 | 0.038 | 0.126 | 0.122 | 0.075 | 0.037 | 0.166 | 0.111 | 0.077 |  |
| 1996 | 0.091 | 0.146 | 0.197 | 0.142 | 0.042 | 0.180 | 0.095 | 0.124 | 0.037 | 0.289 | 0.460 | 0.208 |  |
| 1997 | 0.107 | 0.184 | 0.182 | 0.154 | 0.041 | 0.139 | 0.108 | 0.100 | 0.041 | 0.282 | 0.354 | 0.128 |  |
| 1998 | 0.094 | 0.182 | 0.204 | 0.124 | 0.033 | 0.127 | 0.143 | 0.061 | 0.033 | 0.319 |  | 0.114 |  |
| 1999 | 0.100 | 0.162 | 0.192 | 0.158 | 0.037 | 0.172 | 0.184 | 0.142 | 0.025 | 0.127 | 0.192 | 0.097 |  |
| Avg | 0.094 | 0.147 | 0.161 | 0.125 | 0.039 | 0.144 | 0.130 | 0.096 | 0.032 | 0.243 | 0.279 | 0.150 |  |


|  | Fleet |  |  |  | Fleet |  |  |  | Fleet |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | KR |  |  |  | PG |  |  |  |  |  |  |  |
| Year | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled |
| 1988 | 0.041 | 0.098 | 0.106 | 0.085 |  |  |  |  | 0.080 | 0.128 | 0.122 | 0.118 |
| 1989 | 0.030 | 0.156 | 0.230 | 0.093 |  |  |  |  | 0.057 | 0.169 | 0.179 | 0.157 |
| 1990 | 0.028 | 0.272 | 0.354 | 0.157 |  |  |  |  | 0.120 | 0.244 | 0.240 | 0.232 |
| 1991 | 0.032 | 0.181 | 0.172 | 0.097 |  |  |  |  | 0.046 | 0.153 | 0.211 | 0.156 |
| 1992 | 0.040 | 0.098 | 0.100 | 0.068 |  |  |  |  | 0.106 | 0.123 | 0.134 | 0.126 |
| 1993 | 0.030 | 0.100 | 0.101 | 0.054 |  |  |  |  | 0.053 | 0.131 | 0.123 | 0.115 |
| 1994 | 0.029 | 0.217 | 0.168 | 0.095 | 0.063 | 0.291 | 0.234 | 0.118 | 0.093 | 0.242 | 0.226 | 0.180 |
| 1995 | 0.038 | 0.164 | 0.217 | 0.073 | 0.078 | 0.162 | 0.160 | 0.102 | 0.102 | 0.164 | 0.162 | 0.147 |
| 1996 | 0.041 | 0.234 | 0.239 | 0.099 | 0.055 | 0.222 | 0.222 | 0.101 | 0.084 | 0.242 | 0.236 | 0.223 |
| 1997 | 0.035 | 0.224 | 0.277 | 0.108 | 0.017 | 0.306 | 0.245 | 0.268 | 0.060 | 0.238 | 0.244 | 0.236 |
| 1998 | 0.031 | 0.231 | 0.252 | 0.060 | 0.029 | 0.237 | 0.244 | 0.233 | 0.017 | 0.247 | 0.237 | 0.219 |
| 1999 | 0.030 | 0.220 | 0.236 | 0.086 | 0.052 | 0.232 | 0.241 | 0.238 | 0.099 | 0.243 | 0.246 | 0.237 |
| Avg | 0.034 | 0.183 | 0.204 | 0.090 | 0.049 | 0.242 | 0.224 | 0.177 | 0.076 | 0.194 | 0.197 | 0.179 |

Table 4 con't. Proportions of bigeye tuna in the yellowfin+bigeye by set type (unassociated, log and FAD) and fleet estimated by tree-based regression.

|  | Fleet |  |  |  | Fleet |  |  |  | Fleet |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | SB |  |  |  | SU |  |  |  |  |  |  |  |
| Year | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled |
| 1988 | 0.029 | 0.135 | 0.141 | 0.131 |  |  |  |  | 0.061 | 0.126 | 0.136 | 0.122 |
| 1989 | 0.029 | 0.186 | 0.187 | 0.177 |  |  |  |  | 0.073 | 0.179 | 0.239 | 0.163 |
| 1990 | 0.026 | 0.244 | 0.230 | 0.208 |  |  |  |  | 0.045 | 0.248 | 0.232 | 0.233 |
| 1991 | 0.027 | 0.191 | 0.197 | 0.159 |  |  |  |  | 0.061 | 0.175 | 0.196 | 0.158 |
| 1992 | 0.029 | 0.125 | 0.130 | 0.104 |  |  |  |  | 0.046 | 0.134 | 0.140 | 0.094 |
| 1993 | 0.031 | 0.117 | 0.129 | 0.110 | 0.031 | 0.116 | 0.114 | 0.068 | 0.042 | 0.126 | 0.123 | 0.065 |
| 1994 | 0.033 | 0.236 | 0.244 | 0.197 | 0.025 | 0.222 | 0.222 | 0.133 | 0.039 | 0.254 | 0.266 | 0.101 |
| 1995 | 0.034 | 0.163 | 0.174 | 0.139 |  |  |  |  | 0.048 | 0.155 | 0.180 | 0.071 |
| 1996 | 0.035 | 0.252 | 0.257 | 0.204 |  |  |  |  | 0.071 | 0.236 | 0.235 | 0.117 |
| 1997 | 0.033 | 0.244 | 0.245 | 0.196 |  |  |  |  | 0.044 | 0.253 | 0.261 | 0.123 |
| 1998 | 0.021 | 0.231 | 0.245 | 0.187 |  |  |  |  | 0.031 | 0.231 | 0.234 | 0.051 |
| 1999 | 0.023 | 0.251 | 0.240 | 0.179 |  |  |  |  | 0.033 | 0.247 | 0.257 | 0.148 |
| Avg | 0.029 | 0.198 | 0.202 | 0.166 | 0.028 | 0.169 | 0.168 | 0.100 | 0.050 | 0.197 | 0.208 | 0.121 |


|  | Fleet |  |  |  | Fleet <br> VU |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Unass | Log | FAD | Pooled | Unass | Log | FAD | Pooled |
| 1988 | 0.050 | 0.127 | 0.114 | 0.088 |  |  |  |  |
| 1989 | 0.033 | 0.153 | 0.262 | 0.065 |  |  |  |  |
| 1990 | 0.027 | 0.228 | 0.244 | 0.062 |  |  |  |  |
| 1991 | 0.028 | 0.212 | 0.121 | 0.069 |  |  |  |  |
| 1992 | 0.043 | 0.132 | 0.157 | 0.081 |  |  |  |  |
| 1993 | 0.028 | 0.125 | 0.092 | 0.071 |  |  |  |  |
| 1994 | 0.028 | 0.225 | 0.229 | 0.051 | 0.037 | 0.348 |  |  |
| 1995 | 0.033 | 0.150 | 0.278 | 0.072 | 0.035 | 0.171 | 0.282 | 0.113 |
| 1996 | 0.027 | 0.239 | 0.245 | 0.168 | 0.062 | 0.259 | 0.337 | 0.216 |
| 1997 | 0.033 | 0.251 | 0.234 | 0.152 | 0.034 | 0.259 | 0.231 | 0.157 |
| 1998 | 0.029 | 0.238 | 0.250 | 0.096 | 0.032 | 0.252 | 0.223 | 0.100 |
| 1999 | 0.033 | 0.229 | 0.247 | 0.240 | 0.033 | 0.263 | 0.274 | 0.180 |
| Avg | 0.033 | 0.192 | 0.206 | 0.101 | 0.039 | 0.259 | 0.269 | 0.153 |

Figure 1. Comparison of the spatial distribution of samples from the US observer program (left) and actual fleet distribution (1997-2000, right).


Figure 2. Comparison of the spatial distribution of samples from the US port-sampling program (left) and actual fleet distribution (1988-2000, right).


Figure 3. Comparison of the spatial distribution of the samples from observers on non-US vessels (left) and actual fleet distribution (1997-2000, right).


Figure 4. Comparison of the spatial distribution of the samples from the Japanese port-sampling program (left) and actual fleet distribution (1995-2000, right).



Figure 5. Annual proportions of bigeye tuna in yellowfin+bigeye by unassociated set type and observer sampling (A) and port-sampling of Japanese and US vessels (B). Histogram indicates $\mathbf{2 5}{ }^{\text {th }}$ and $\mathbf{7 5}^{\text {th }}$ percentiles, circle indicates the median.


Observer sampling (all vessels) - BET percent in YFT:BET, unass. sets, 1988-2000, $n=143$


Figure 6. Annual proportions of bigeye tuna in yellowfin+bigeye by associated set type and sampling program (observer and portsampling). Histogram indicates $25^{\text {th }}$ and $75^{\text {th }}$ percentiles, circle indicates the median.
A) Observer sampling (Japanese vessels), B) Japanese port-sampling, C) Observer sampling (US vessels) and D) US port-sampling





Figure 10. Spatial distribution of bigeye catch by set type (white - unassociated, black - log, grey - drifting FAD) for purse seine fleets in 1999. Bigeye catch was estimated by tree-based regression.
$\square$ Unass Log
FAD




Figure 7 . Generalized additive model (GAM) derived effects of predictor variables, school type (A), year (B), month (C), fleet (D), latitude (E) and longitude ( $F$ ) on the composition of bigeye in yellowfin+bigeye estimates. Relative density of data points is shown on the $\mathbf{x}$-axis (rug-plot). All variables were categorical except latitude and longitude which were continuous.






$60 \begin{array}{lll} \\ & 180 \\ \text { lon5 }\end{array}$ lon5

Figure 8. Pruned tree for the proportion of bigeye in the yellowfin+bigeye estimates for the purse seine fishery in the western and central Pacific Ocean. Initial 30 nodes illustrated.


Figure 9. Annual bigeye catch by all purse seine fleets in the WCPO estimated from two methods (extrapolation of US port-sampling data and tree-based regression).


